4.2 — Difference-in-Difference Models ECON 480 • Econometrics • Fall 2021 Ryan Safner Assistant Professor of Economics ✓ safner@hood.edu

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Outline

<u>Difference-in-Difference Models</u>

- Example I: HOPE in Georgia
- **Generalizing DND Models**
- Example II: "The" Card-Kreuger Minimum Wage Study



Clever Research Designs Identify Causality



Again, **this toolkit** of research designs to **identify causal effects** is the economist's **comparative advantage** that firms and governments want!



Natural Experiments



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- **Example**: how do States that implement policy *X* see changes in *Y*
 - **Treatment**: States that implement X
 - $\circ~{\bf Control}:$ States that did not implement X

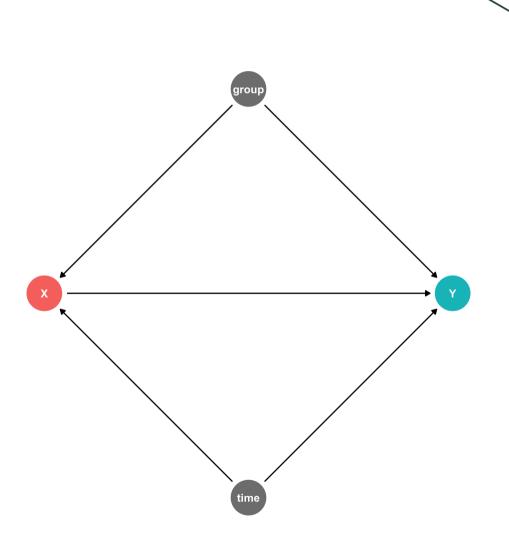


- Often, we want to examine the consequences of a change, such as a law or policy intervention
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- Find the *difference* between treatment & control groups *in* their *differences* before and after the treatment period





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 - \circ **Control**: States that did not implement X
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• The difference-in-difference model (aka "diff-in-diff" or "DND") identifies treatment effect by differencing the difference pre- and post-treatment values of Y between treatment and control groups

$$\hat{Y}_{it} = \beta_0 + \beta_1 \operatorname{Treated}_i + \beta_2 \operatorname{After}_t + \beta_3 (\operatorname{Treated}_i \times \operatorname{After}_t) + u_{it}$$

• *Treated*_i = $\begin{cases} 1 \text{ if } i \text{ is in treatment group} \\ 0 \text{ if } i \text{ is not in treatment group} \end{cases}$ $After_t = \begin{cases} 1 \text{ if } t \text{ is after treatment period} \\ 0 \text{ if } t \text{ is before treatment period} \end{cases}$

	Control	Treatment	Group Diff (ΔY_i)
Before	β_0	$\beta_0 + \beta_1$	β_1
After	$\beta_0 + \beta_2$	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_1 + \beta_3$
Time Diff (ΔY_t)	β_2	$\beta_2 + \beta_3$	Diff-in-diff $\Delta_i \Delta_t : \beta_3$

Silly Example: Hot Dogs



Is there a discount when you get cheese *and* chili?

price	cheese	chili
<qp[></qp[>	<qpf></qpf>	<qpf></qpf>
2.00	0	0
2.35	1	0
2.35	0	1
2.70	1	1
4 rows		

```
lm(price ~ cheese + chili + cheese*chili,
    data = hotdogs) %>%
    tidy()
```

term	estimate
<chr></chr>	<dpl></dpl>
(Intercept)	2.00
cheese	0.35
chili	0.35
cheese:chili	0.00
4 rows	

Silly Example: Hot Dogs



Is there a discount when you get cheese and chili?

	No Cheese	Cheese	Cheese Diff
No Chili	\$2.00	\$2.35	\$0.35
Chili	\$2.35	\$2.70	\$0.35
Chili Diff	\$0.35	\$0.35	\$0.00 (Diff-in-diff)

• Diff-n-diff is just a model with an interaction term between two dummies!

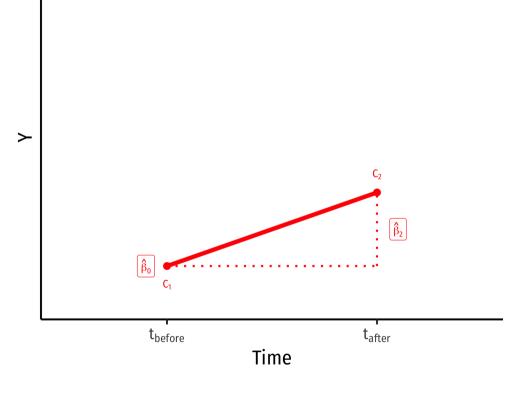
```
lm(price ~ cheese + chili + cheese*chili,
    data = hotdogs) %>%
    tidy()
```

term	estimate
<chr></chr>	<qpf></qpf>
(Intercept)	2.00
cheese	0.35
chili	0.35
cheese:chili	0.00
4 rows	

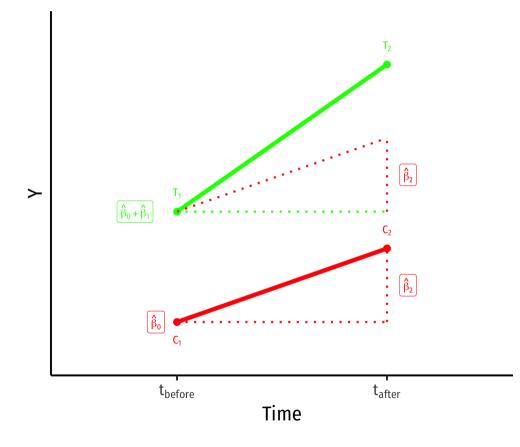




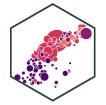
- Control group (Treated_i = 0)
- $\hat{\beta}_0$: value of *Y* for **control** group **before** treatment
- $\hat{\beta}_2$: time *difference* (for **control** group)

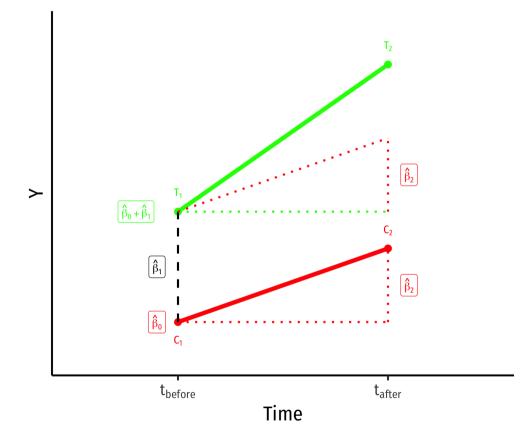






- Control group (Treated_i = 0)
- $\hat{\beta}_0$: value of *Y* for **control** group **before** treatment
- $\hat{\beta}_2$: time *difference* (for **control** group)
- Treatment group (Treated_i = 1)

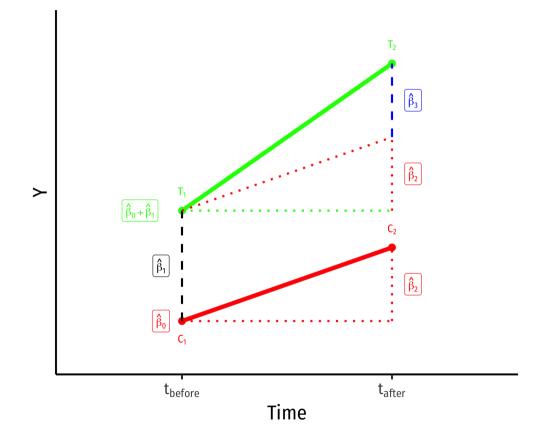




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- $\hat{\beta}_2$: time *difference* (for **control** group)
- Treatment group (Treated_i = 1)
- $\hat{\beta}_1$: *difference* between groups **before** treatment



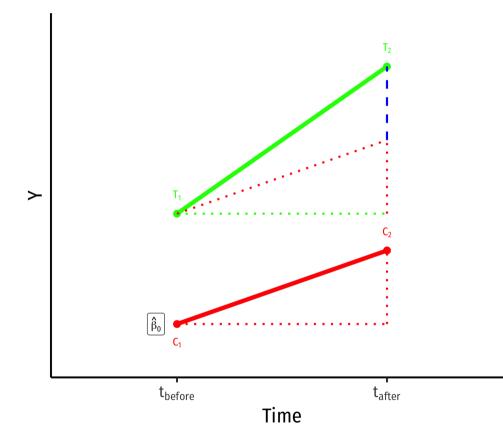
 $\hat{Y}_{it} = \beta_0 + \beta_1 \text{Treated}_i + \beta_2 \text{After}_t + \beta_3 (\text{Treated}_i \times \text{After}_t) + u_{it}$



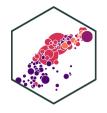
- Control group (Treated_i = 0)
- $\hat{\beta}_0$: value of *Y* for **control** group **before** treatment
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- $\hat{\beta}_1$: *difference* between groups **before** treatment

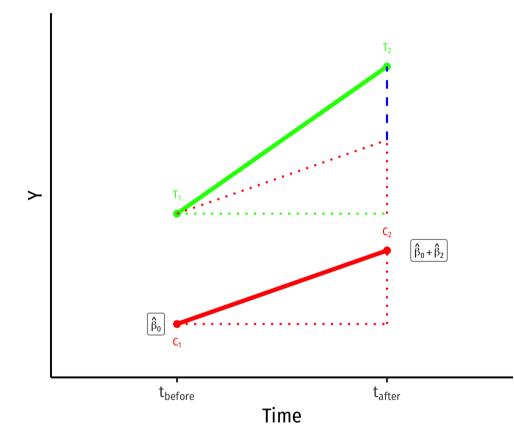
• $\hat{\beta}_3$: difference-in-difference (treatment effect)

 $\hat{Y}_{it} = \beta_0 + \beta_1 \text{Treated}_i + \beta_2 \text{After}_t + \beta_3 (\text{Treated}_i \times \text{After}_t) + u_{it}$



• \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$

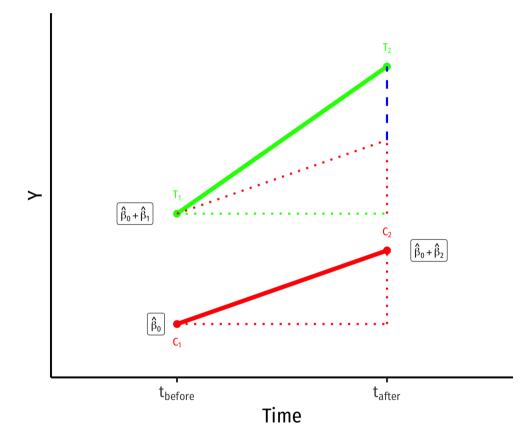




- \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$
- \bar{Y}_i for **Control** group **after**: $\hat{\beta}_0 + \hat{\beta}_2$

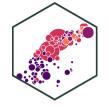


 $\hat{Y}_{it} = \beta_0 + \beta_1 \text{Treated}_i + \beta_2 \text{After}_t + \beta_3 (\text{Treated}_i \times \text{After}_t) + u_{it}$

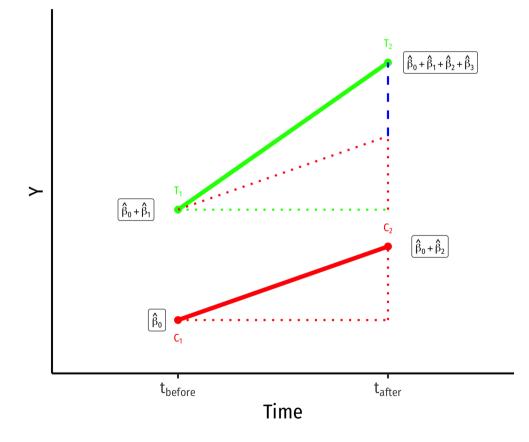


- \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$
- \bar{Y}_i for **Control** group **after**: $\hat{\beta}_0 + \hat{\beta}_2$

• \bar{Y}_i for **Treatment** group **before**: $\hat{\beta}_0 + \hat{\beta}_1$



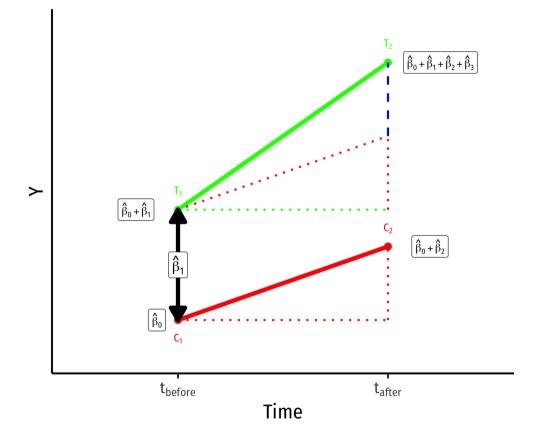




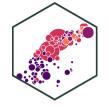
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- \bar{Y}_i for **Treatment** group **before**: $\hat{\beta}_0 + \hat{\beta}_1$
- \bar{Y}_i for **Treatment** group **after**: $\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3$

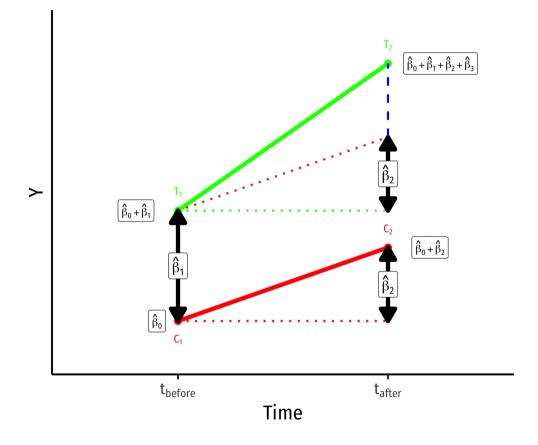




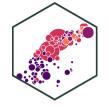


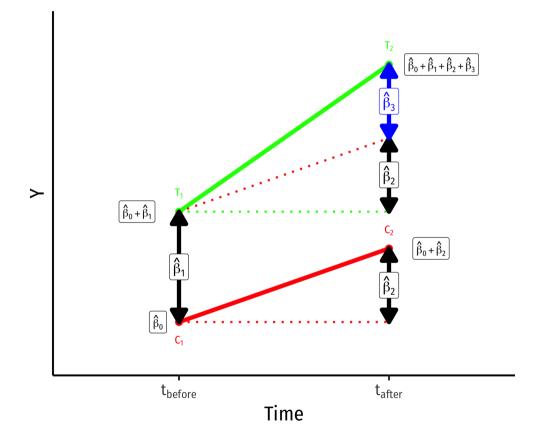
- \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$
- \bar{Y}_i for **Control** group **after**: $\hat{\beta}_0 + \hat{\beta}_2$
- \bar{Y}_i for **Treatment** group **before**: $\hat{\beta}_0 + \hat{\beta}_1$
- \bar{Y}_i for **Treatment** group **after**: $\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3$
- Group Difference (before): $\hat{\beta_1}$





- \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$
- \bar{Y}_i for **Control** group **after**: $\hat{\beta}_0 + \hat{\beta}_2$
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- \bar{Y}_i for **Treatment** group **after**: $\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3$
- Group Difference (before): $\hat{\beta}_1$
- Time Difference: $\hat{\beta}_2$





- \bar{Y}_i for **Control** group **before**: $\hat{\beta}_0$
- \bar{Y}_i for **Control** group **after**: $\hat{\beta}_0 + \hat{\beta}_2$
- \bar{Y}_i for **Treatment** group **before**: $\hat{\beta}_0 + \hat{\beta}_1$
- \bar{Y}_i for **Treatment** group **after**: $\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3$
- Group Difference (before): $\hat{\beta_1}$
- Time Difference: $\hat{\beta}_2$
- **Difference-in-difference**: $\hat{\beta}_3$ (treatment effect)

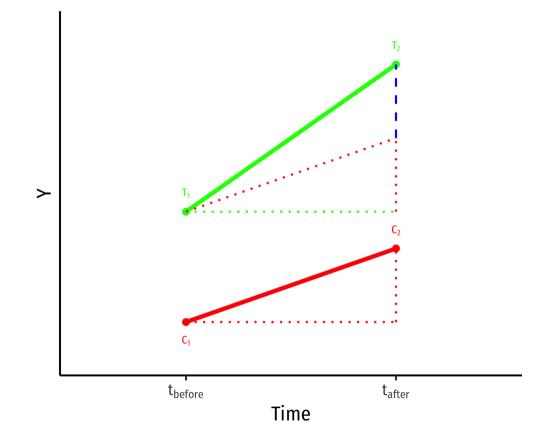
Comparing Group Means (Again)



	Control	Treatment	Group Diff (ΔY_i)
Before	β_0	$\beta_0 + \beta_1$	β_1
After	$\beta_0 + \beta_2$	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_1 + \beta_3$
Time Diff (ΔY_t)	β_2	$\beta_2 + \beta_3$	Diff-in-diff $\Delta_i \Delta_t : \beta_3$

Key Assumption: Counterfactual



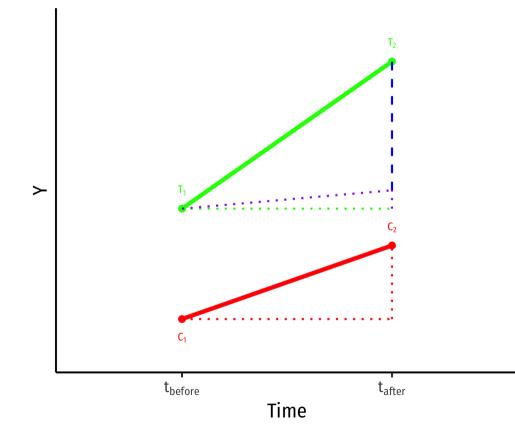


- Key assumption for DND: **time trends** (for treatment and control) are **parallel**
- Treatment and control groups assumed to be identical over time on average, except for treatment
- Counterfactual: if the treatment group had not recieved treatment, it would have changed identically over time as the control group $(\hat{\beta}_2)$

Key Assumption: Counterfactual



 $\hat{Y}_{it} = \beta_0 + \beta_1 \text{Treated}_i + \beta_2 \text{After}_t + \beta_3 (\text{Treated}_i \times \text{After}_t) + u_{it}$



• If the time-trends would have been *different*, a **biased** measure of the treatment effect $(\hat{\beta}_3)!$



Example I: HOPE in Georgia

Diff-in-Diff Example I

Example: In 1993 Georgia initiated a HOPE scholarship program to let state residents with at least a B average in high school attend public college in Georgia for free. Did it increase college enrollment?

- Micro-level data on 4,291 young individuals
- InCollege_{*it*} = $\begin{cases} 1 \text{ if } i \text{ is in college during year } t \\ 0 \text{ if } i \text{ is not in college during year } t \end{cases}$
- Georgia_i = $\begin{cases} 1 \text{ if } i \text{ is a Georgia resident} \\ 0 \text{ if } i \text{ is not a Georgia resident} \end{cases}$
- After_t = $\begin{cases} 1 \text{ if } t \text{ is after } 1992 \\ 0 \text{ if } t \text{ is after } 1992 \end{cases}$

Diff-in-Diff Example II

- We can use a DND model to measure the effect of HOPE scholarship on enrollments
- Georgia and nearby States, if not for HOPE, changes should be the same over time
- Treatment period: after 1992
- Treatment: Georgia
- Differences-in-differences:

 $\Delta_i \Delta_t Enrolled = (GA_{after} - GA_{before}) - (neighbors_{after} - neighbors_{before})$

• Regression equation:

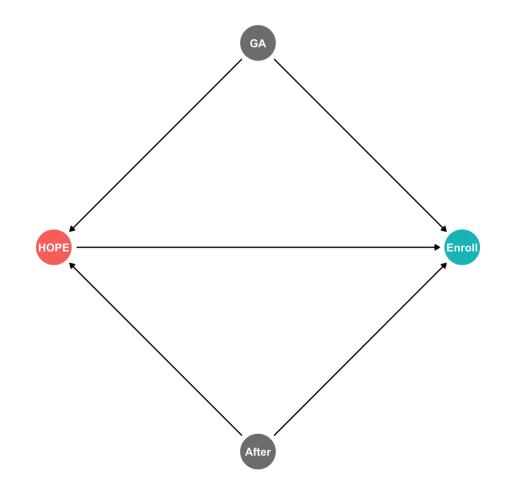
 $\widehat{\text{Enrolled}}_{it} = \beta_0 + \beta_1 \operatorname{Georgia}_i + \beta_2 \operatorname{After}_t + \beta_3 (\operatorname{Georgia}_i \times \operatorname{After}_t)$

Example: Data

StateCode	A Year	Weight	Age18	LowIncome	InCollege	After	Georgia	AfterGeorgia
<fct></fct>	<dbl×fct></dbl×fct>	<qp[></qp[>	<qpf></qpf>	<dpl></dpl>	<qpf></qpf>	<qpf></qpf>	<dpl></dpl>	<dpl></dpl>
56	19 89	1396	0	1	1	0	0	0
56	19 89	1080	0	NA	1	0	0	0
56	18 89	924	1	1	1	0	0	0
56	19 89	891	0	0	1	0	0	0
56	19 89	1395	0	NA	0	0	0	0
56	18 89	1106	1	1	1	0	0	0
56	19 89	965	0	NA	0	0	0	0
56	18 89	958	1	NA	0	0	0	0
56	19 89	1006	0	NA	0	0	0	0
56	19 89	1183	0	1	1	0	0	0
1-10 of 4,291	rows 1-10 of	f 11 colum	ns			Previo	us 1 2 3	4 5 6 430Next

Example: Data





Example: Regression



DND_reg <- lm(InCollege ~ Georgia + After + Georgia*After, data = hope)
DND_reg %>% tidy()

term	estimate	std.error	statistic	p.value
<chr></chr>	<dpf></dpf>	<qpf></qpf>	<qpf></qpf>	<qpf></qpf>
(Intercept)	0.405782652	0.01092390	37.1463182	4.221545e-262
Georgia	-0.105236204	0.03778114	-2.7854165	5.369384e-03
After	-0.004459609	0.01585224	-0.2813235	7.784758e-01
Georgia:After	0.089329828	0.04889329	1.8270364	6.776378e-02
4 rows				

 $\widehat{\text{Enrolled}}_{it} = 0.406 - 0.105 \operatorname{Georgia}_i - 0.004 \operatorname{After}_t + 0.089 (\operatorname{Georgia}_i \times \operatorname{After}_t)$

Example: Interpretting the Regression



 $\widehat{\text{Enrolled}}_{it} = 0.406 - 0.105 \operatorname{Georgia}_i - 0.004 \operatorname{After}_t + 0.089 (\operatorname{Georgia}_i \times \operatorname{After}_t)$

- β_0 : A **non-Georgian before** 1992 was 40.6% likely to be a college student
- β_1 : **Georgians before** 1992 were 10.5% less likely to be college students than neighboring states
- β_2 : After 1992, non-Georgians are 0.4% less likely to be college students
- β_3 : After 1992, Georgians are 8.9% more likely to enroll in colleges than neighboring states
- Treatment effect: HOPE increased enrollment likelihood by 8.9%

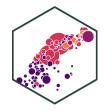
Example: Comparing Group Means

 $\widehat{\text{Enrolled}}_{it} = 0.406 - 0.105 \operatorname{Georgia}_i - 0.004 \operatorname{After}_t + 0.089 (\operatorname{Georgia}_i \times \operatorname{After}_t)$

- A group mean for a dummy Y is E[Y = 1], i.e. the probability a student is enrolled:
- Non-Georgian enrollment probability pre-1992: $\beta_0 = 0.406$
- Georgian enrollment probability pre-1992: $\beta_0 + \beta_1 = 0.406 0.105 = 0.301$
- Non-Georgian enrollment probability post-1992: $\beta_0 + \beta_2 = 0.406 0.004 = 0.402$
- Georgian enrollment probability post-1992:

 $\beta_0 + \beta_1 + \beta_2 + \beta_3 = 0.406 - 0.105 - 0.004 + 0.089 = 0.386$

Example: Comparing Group Means in R





	prob
	<dpf></dpf>
	0.4057827
1 row	

Example: Comparing Group Means in R II





prob	
<dbl></dbl>	
0.3005464	

Example: Diff-in-Diff Summary



 $\widehat{\text{Enrolled}}_{it} = 0.406 - 0.105 \operatorname{Georgia}_i - 0.004 \operatorname{After}_t + 0.089 (\operatorname{Georgia}_i \times \operatorname{After}_t)$

	Neighbors	Georgia	Group Diff (ΔY_i)
Before	0.406	0.301	-0.105
After	0.402	0.386	0.016
Time Diff (ΔY_t)	-0.004	0.085	Diff-in-diff: 0.089

 $\Delta_i \Delta_t Enrolled = (GA_{after} - GA_{before}) - (neighbors_{after} - neighbors_{before})$ = (0.386 - 0.301) - (0.402 - 0.406) = (0.085) - (-0.004) = 0.089

Diff-in-Diff Summary & Data



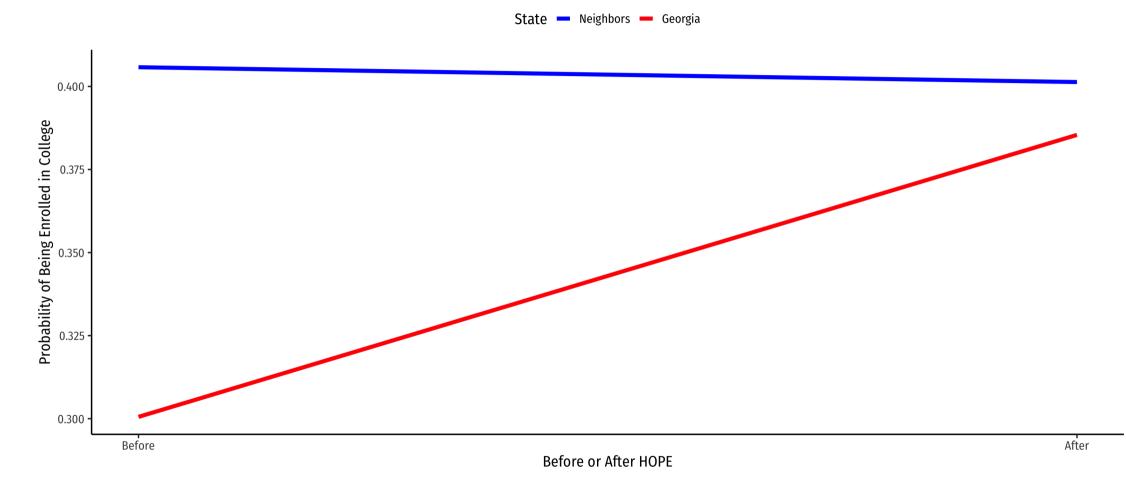
TABLE 2DIFFERENCE-IN-DIFFERENCESSHARE OF 18-19-YEAR-OLDS ATTENDING COLLEGEOCTOBER CPS, 1989-97

	Before 1993	1993 and After	Difference
Georgia	0.300	0.378	0.078
Rest of Southeastern States	0.415	0.414	-0.001
Difference	0.115	0.036	0.079

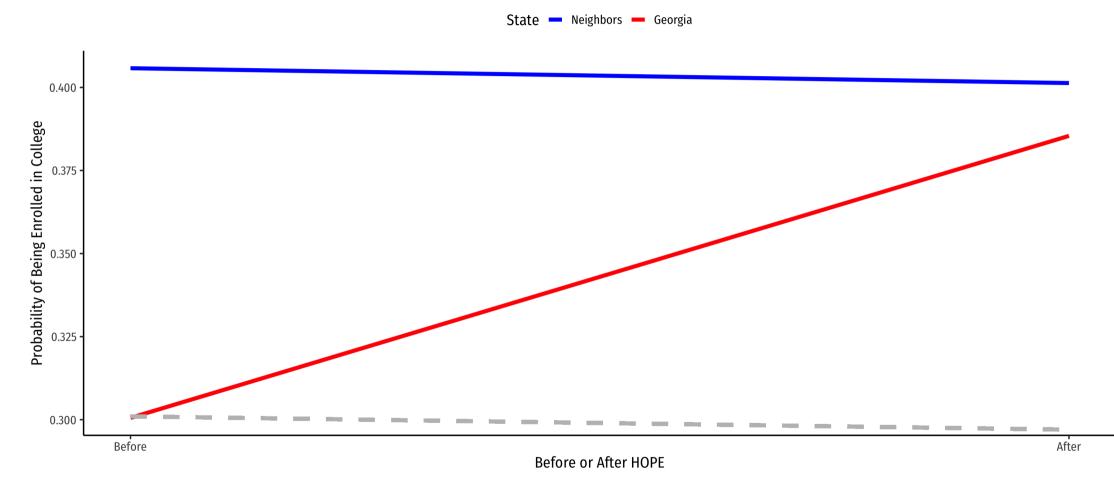
Note: Means are weighted by CPS sample weights. The Southeastern states are defined in the note to Table 1.

Dynarski, Susan, 1999, "Hope for Whom? Financial Aid for the Middle Class and its Impact on College Attendance," *National Tax Journal* 53(3): 629-661

Example: Diff-in-Diff Graph



Example: Diff-in-Diff Graph (& Counterfactual)





Generalizing DND Models

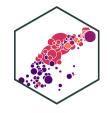
Generalizing DND Models

• DND can be **generalized** with a **two-way fixed effects** model:

 $\widehat{Y_{it}} = \beta_1(\text{Treated}_i \times \text{After}_t) + \alpha_i + \theta_t + \nu_{it}$

- α_i : group fixed effects (treatments/control groups)
- θ_t : **time fixed effects** (pre/post treatment)
- β_1 : diff-in-diff (interaction effect, β_3 from before)
- Flexibility: *many* periods (not just before/after), *many* different treatment(s)/groups, and treatment(s) can occur at different times to different units (so long as some do not get treated)
- Can also add control variables that vary within units and over time

 $\widehat{Y_{it}} = \beta_1 (\text{Treated}_i \times \text{After}_t) + \beta_2 X_{it} + \dots + \alpha_i + \theta_t + \nu_{it}$



Our Example, Generalized I



$$\widehat{\text{Enrolled}}_{it} = \beta_1 \left(\operatorname{Georgia}_i \times \operatorname{After}_t \right) + \alpha_i + \theta_t + \theta_t$$

- StateCode is a variable for the State \implies create State fixed effect
- Year is a variable for the year \implies create year fixed effect

Our Example, Generalized II

• Using LSDV method...

term	estimate	std.error	statistic	p.value
<chr></chr>	<qpf></qpf>	<dpl></dpl>	<qpf></qpf>	<qpf><qpf><</qpf></qpf>
(Intercept)	0.418057478	0.02261133	18.4888517	1.734550e-73
Georgia	-0.141501255	0.03936119	-3.5949436	3.281224e-04
After	0.075340594	0.03128021	2.4085706	1.605717e-02
factor(StateCode)57	-0.014181112	0.02739708	-0.5176140	6.047544e-01
factor(StateCode)58	NA	NA	NA	NA
factor(StateCode)59	-0.062378540	0.01954266	-3.1919172	1.423556e-03
factor(StateCode)62	-0.132650271	0.02806143	-4.7271383	2.350298e-06
factor(StateCode)63	-0.005103868	0.02627780	-0.1942274	8.460071e-01
factor(Year)90	0.046608845	0.02833625	1.6448486	1.000745e-01
factor(Year)91	0.032275789	0.02856877	1.1297577	2.586417e-01
1-10 of 17 rows			Previous	1 2 Next

• By de-meaning data, using fixest

library(fixest)

estimate statistic term std.error p.value <chr> <dbl> <dbl> <dbl> <dbl> Georgia:After 16.19978 1.633762e-05 0.0914202 0.005643298 1 row

 $\widehat{\text{InCollege}}_{it} = 0.091 (\operatorname{Georgia}_i \times \operatorname{After}_{it}) + \alpha_i + \theta_t$



Our Example, Generalized, with Controls II

• Using LSDV method...

term	estimate	std.error	statistic	p.value
<chr></chr>	<qpf><</qpf>	<dpl></dpl>	<qpf></qpf>	<qpf><qpf><</qpf></qpf>
(Intercept)	0.735574222	0.02990710	24.5953037	1.155308e-121
Georgia	-0.108940276	0.04765262	-2.2861342	2.231699e-02
After	-0.005753553	0.03737027	-0.1539607	8.776512e-01
factor(StateCode)57	-0.043406073	0.03047696	-1.4242257	1.544869e-01
factor(StateCode)58	NA	NA	NA	NA
factor(StateCode)59	-0.053175645	0.02306160	-2.3058092	2.119033e-02
factor(StateCode)62	-0.116104615	0.03283310	-3.5362060	4.121675e-04
factor(StateCode)63	0.007389866	0.03056444	0.2417799	8.089675e-01
factor(Year)90	0.039364315	0.03326291	1.1834297	2.367342e-01
factor(Year)91	0.029227969	0.03347850	0.8730370	3.827140e-01
1-10 of 19 rows			Previous	1 2 Next

• By de-meaning data, using fixest

library(fixest)

estimate statistic term std.error p.value <dbl> <dbl> <dbl> <chr> <dbl> Black -0.09398715 0.01273233 -7.381769 0.0007172767 LowIncome -0.30172426 0.03066188 -9.840369 0.0001846526 Georgia:After 0.02343679 0.01281838 1.828374 0.1270331132 3 rows

 $\widehat{\text{InCollege}}_{it} = 0.023 (\operatorname{Georgia}_i \times \operatorname{After}_{it}) - 0.094 \operatorname{Black}_{it} - 0.302 \operatorname{LowIncome}_{it}$

Our Example, Generalized, with Controls II

\wedge	

	(1)	(2)	(3)
Intercept	0.4058 ***		
	(0.0109)		
Georgia	-0.1052 **		
	(0.0378)		
After	-0.0045		
	(0.0159)		
Georgia x After	0.0893	0.0914 ***	0.0234
	(0.0489)	(0.0056)	(0.0128)
Low Income			-0.3017 ***
			(0.0307)
Black			-0.0940 ***
			(0.0127)
Fixed Effects	None	State & Year	State & Year
Ν	4291	4291	2967
R-Squared	0.0019	0.0115	0.1036
SER	0.4893	0.4875	0.4731

*** p < 0.001; ** p < 0.01; * p < 0.05.

Our Example, Generalized, with Controls II



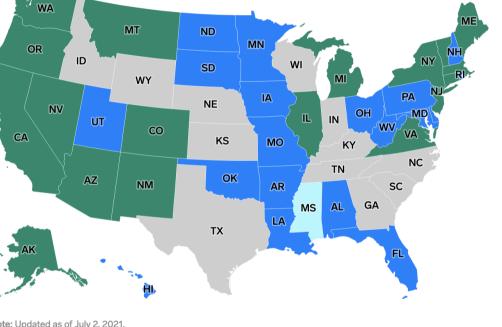
TABLE 3 COLLEGE ATTENDANCE OF 18–19–YEAR–OLDS OCTOBER CPS, 1989–97 CONTROL GROUP: SOUTHEASTERN STATES				
	(1) Difference–in– Differences	(2) Add Covariates	(3) Add Local Economic Conditions Controls	
After*Georgia	0.079 (0.029)	0.075 (0.030)	0.070 (0.030)	
Georgia	-0.115 (0.023)	-0.100 (0.019)	-0.097 (0.018)	
After	-0.001 (0.018)			
Age 18		-0.042 (0.014)	-0.042 (0.016)	
Metro Resident		0.042 (0.016)	0.042 (0.015)	
Black		-0.134 (0.014)	-0.133 (0.015)	
State Unemployment Rate			0.005 (0.007)	
Year Dummies R² N	0.003 6,811	Yes 0.023 6,811	Yes 0.023 6,811	

Note: Regressions are weighted by CPS sample weights. Standard errors are adjusted for heteroskedasticity and correlation within state-year cells. The Southeastern states are defined in the note to Table 1.

Intuition behind DND

- Diff-in-diff models are the quintessential example of exploiting natural experiments
- A major change at a point in time (change in law, a natural disaster, political crisis) separates groups where one is affected and another is not---identifies the effect of the change (treatment)
- One of the cleanest and clearest causal identification strategies









Example II: "The" Card-Kreuger Minimum Wage Study

Example: "The" Card-Kreuger Minimum Wage Study I

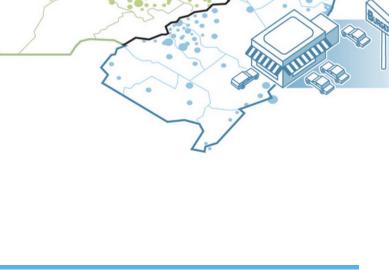


Example: *The* controversial minimum wage study, Card & Kreuger (1994) is a quintessential (and clever) diff-in-diff.

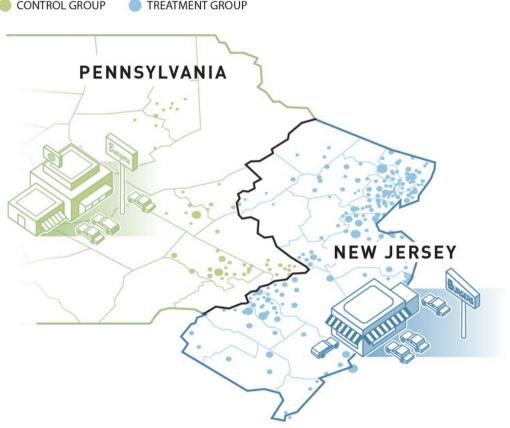
Card, David, Krueger, Alan B, (1994), "Minimum Wages and Employment: A Case Study of the Fast-Food Industry in New Jersey and Pennsylvania," *American Economic Review* 84 (4): 772–793

Card & Kreuger (1994): Background I

- Card & Kreuger (1994) compare employment in fast food restaurants on New Jersey and Pennsylvania sides of border between February and November 1992.
- Pennsylvania & New Jersey both had a minimum wage of \$4.25 before February 1992
- In February 1992, New Jersey raised minimum wage from \$4.25 to \$5.05

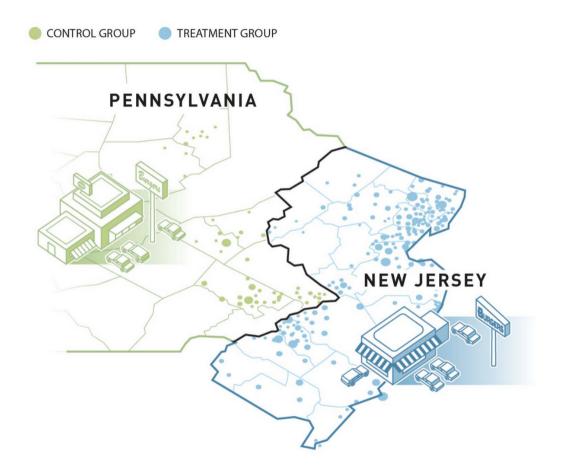






Card & Kreuger (1994): Background II

- If we look only at New Jersey before & after change:
 - Omitted variable bias:
 - macroeconomic variables (there's a recession!), weather, etc.
 - Including PA as a control will control for these time-varying effects if they are national trends
- Surveyed 400 fast food restaurants on each side of the border, before & after min wage increase





Card & Kreuger (1994): Comparisons



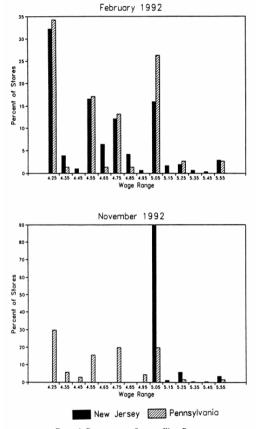


FIGURE 1. DISTRIBUTION OF STARTING WAGE RATES

Card & Kreuger (1994): Summary I

TABLE 1-SAMPLE DESIGN AND RESPONSE RATES

		Sto	ores in:
	All	NJ	PA
Wave 1, February 15–March 4, 1992:			
Number of stores in sample frame: ^a	473	364	109
Number of refusals:	63	33	30
Number interviewed:	410	331	79
Response rate (percentage):	86.7	90.9	72.5
Wave 2, November 5 – December 31, 1992:			
Number of stores in sample frame:	410	331	79
Number closed:	6	5	1
Number under rennovation:	2	2	0
Number temporarily closed: ^b	2	2	0
Number of refusals:	1	1	0
Number interviewed: ^c	399	321	78

Card & Kreuger (1994): Summary II



	Stores in:		
Variable	NJ	PA	
1. Distribution of Store Types (per			
a. Burger King	41.1	44.3	
b. KFC	20.5	15.2	
· ·			
b. KFC	20.5	15.2	

Card & Kreuger (1994): Model



 $\widehat{\text{Employment}_{it}} = \beta_0 + \beta_1 \operatorname{NJ}_i + \beta_2 \operatorname{After}_t + \beta_3 (\operatorname{NJ}_i \times After_t)$

- PA Before: β_0
- PA After: $\beta_0+\beta_2$
- NJ Before: $\beta_0 + \beta_1$
- NJ After: $\beta_0 + \beta_1 + \beta_2 + \beta_3$
- Diff-in-diff: $(NJ_{after} NJ_{before}) (PA_{after} PA_{before})$

	PA	NJ	Group Diff (ΔY_i)
Before	β_0	$\beta_0 + \beta_1$	β_1
After	$\beta_0 + \beta_2$	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_1 + \beta_3$
Time Diff (ΔY_t)	β_2	$\beta_2 + \beta_3$	Diff-in-diff $\Delta_i \Delta_t : eta_3$

Card & Kreuger (1994): Results



	Stores by state		
Variable	PA (i)	NJ (ii)	Difference, NJ – PA (iii)
 FTE employment before,	23.33	20.44	-2.89
all available observations	(1.35)	(0.51)	(1.44)
FTE employment after,	21.17	21.03	-0.14
all available observations	(0.94)	(0.52)	(1.07)
 Change in mean FTE	-2.16	0.59	2.76
employment	(1.25)	(0.54)	(1.36)